Leverage Mendelian Randomization to Learn Meaningful Representations (LMR×2)

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#### Outline

What is MR?

Summary-data MR

Mechanistic heterogeneity

## What is MR?

#### Wikipedia definition:

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MR = Use genetic variation as instrumental variables.

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Folk definition:

MR = Use genetic variation as instrumental variables.

A more informative definition:

MR = Base causal inference on randomness in Mendelian inheritance.

#### Heredity as a natural experiment

#### **Autosomal Dominant Inheritance Pattern**



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# Surging popularity of MR



► Applications of MR are fueled by the increasing availability of GWAS datasets.<sup>1</sup>

 $<sup>^{1}\</sup>textsc{Data}$  are obtained from Web of Science (https://www.webofknowledge.com/).

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A well understood pathway of heart disease



#### Basic idea

People who inherited certain alleles of *rs17238484* and *rs12916* have **naturally** higher concentration of LDL cholesterol.

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Must assume 3 core IV assumptions  $\implies$  Partial identification (1) Relevance:  $Z \not\perp X$ .

- (2) Exogeneity (natural experiment):  $Z \perp U$ .
- **3** Exclusion restriction: *Z* has no direct effect on *Y*.

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(3) Exclusion restriction: Z has no direct effect on Y.

#### Plus 1 extra assumption $\implies$ Point identification

Could be linearity, monotonicity (Angrist, Imbens & Rubin, 1996), or homogeneity (Hernán & Robins, 2006; Wang & Tchetgen Tchetgen, 2018).

## Basic idea: division



#### The Wald estimator

Causal effect of X on Y (
$$\beta_0$$
) =  $\frac{\text{Causal effect of } Z \text{ on } Y (\Gamma = \gamma \cdot \beta_0)}{\text{Causal effect of } Z \text{ on } X (\gamma)}$ .

#### Basic idea: division



The Wald estimator

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Heuristic: Linear structural equation model

$$X = \gamma Z + \eta_X U + E_X,$$
  

$$Y = \beta_0 X + \eta_Y U + E_Y$$
  

$$= (\beta_0 \gamma) Z + \underbrace{f(U, E_X, E_Y)}_{\text{index}}$$

independent of Z

#### Example: Causal effect of LDL-cholesterol



#### A main challenge to MR

#### Violation of exclusion restriction due to pleiotropy (multiple functions of genes)

 $<sup>^{2}</sup>$ Swerdlow, D. I., et al. "HMG-coenzyme A reductase inhibition, type 2 diabetes, and bodyweight: evidence from genetic analysis and randomised trials." *Lancet* (2015).

<sup>&</sup>lt;sup>3</sup>Boyle, E. et al. (2017). "An expanded view of complex traits: from polygenic to omnigenic". *Cell* 169, p1177–1186.

## A main challenge to MR

#### Violation of exclusion restriction due to pleiotropy (multiple functions of genes)

Example: *HMGCR* is associated with body weight<sup>2</sup>



#### Recent genetic studies show that pleiotropy is indeed wide-spread.<sup>3</sup>

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- $1.\ <50\%$  of the calipers are broken (Kang et al., 2016); or
- 2. Rusty readings are balanced around the truth (Bowden et al., 2015).

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#### Remaining issues

- 1. Both situations are common in MR.
- 2. Need to deal with many weak instruments.

### Three-sample summary-data MR

- Sample 1: Select genetic variants associated with the hypothesized cause (LDL-C in the previous example; epidemiologists call this exposure).
- Sample 2: Obtain the GWAS summary data (γ̂<sub>j</sub>, σ<sub>Xj</sub>), j = 1,..., p for the gene-exposure associations.
- Sample 3: Obtain the GWAS summary data  $(\hat{\Gamma}_j, \sigma_{Yj}), j = 1, ..., p$  for the gene-outcome associations.

This is crucial for eliminating selection bias and the dependence between  $\hat{\gamma}_j$  and  $\hat{\Gamma}_j$ .

#### Assumptions

Assumption 1: Measurement error model

$$\begin{pmatrix} \hat{\boldsymbol{\gamma}} \\ \hat{\boldsymbol{\Gamma}} \end{pmatrix} \sim \mathrm{N} \left( \begin{pmatrix} \boldsymbol{\gamma} \\ \boldsymbol{\Gamma} \end{pmatrix}, \begin{array}{c} \left( \boldsymbol{\Sigma}_{\boldsymbol{X}} & \boldsymbol{0} \\ \boldsymbol{0} & \boldsymbol{\Sigma}_{\boldsymbol{Y}} \end{array} \right) \right), \quad \boldsymbol{\Sigma}_{\boldsymbol{X}} = \mathrm{diag}(\sigma_{\boldsymbol{X}1}^2, \dots, \sigma_{\boldsymbol{X}\rho}^2), \\ \boldsymbol{\Sigma}_{\boldsymbol{Y}} = \mathrm{diag}(\sigma_{\boldsymbol{Y}1}^2, \dots, \sigma_{\boldsymbol{Y}\rho}^2).$$

#### Assumption 2: Random rusty calipers

The causal effect  $\beta$  satisfies  $\Gamma \approx \beta_0 \gamma$ . Specifically, let  $\alpha_j = \Gamma_j - \beta \gamma_j$ . Then we assume

- ▶ InSIDE (Instrument Strength Independent of Direct Effect):  $\alpha_j$  is independent of  $\gamma_j$ ;
- Most  $\alpha_j \stackrel{\text{i.i.d.}}{\sim} N(0, \tau^2)$ , but a few  $|\alpha_j|$  might be very large.

These assumptions are based on extensive exploratory data analyses.

• Define standardized residual: 
$$t_j(\beta, \tau^2) = \frac{\ddot{\Gamma}_j - \beta \hat{\gamma}_j}{\sqrt{1 + \beta^2 \sigma_{Xj}^2 + \tau^2 \sigma_{Yj}^2}}$$
.

<sup>&</sup>lt;sup>4</sup>Zhao, Q. et al. (2019). "Powerful three-sample genome-wide design and robust statistical inference in summary-data Mendelian randomization". *International Journal of Epidemiology*, 48(5):1478-1492.

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For some robust loss  $\rho$  (let  $\psi = \rho'$ ), the RAPS equations are

$$\psi_1^{(\rho)}(\beta,\tau^2) = \sum_{j=1}^{\rho} \left(\frac{\partial}{\partial\beta} t_j\right) \cdot \psi(t_j),$$
  
$$\psi_2^{(\rho)}(\beta,\tau^2) = \sum_{j=1}^{\rho} t_j \cdot \psi(t_j) - \mathbb{E}[T\psi(T)], \text{ for } T \sim N(0,1).$$

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Roughly speaking, the first equation means that

$$\sum_{j=1}^{p} \begin{pmatrix} \text{Estimated quality} \\ \text{of instrument } j \end{pmatrix} \cdot \begin{pmatrix} \text{Estimated error} \\ \text{of instrument } j \end{pmatrix} = 0.$$

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Estimated quality of the instruments can be improved by empirical Bayes, which works really well with many weak instruments.<sup>4</sup>

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## Motivating example: BMI and type 2 diabetes (T2D)



## Two scenarios of mechanistic heterogeneity



(a) Scenario 1: Multiple pathways of horizontal pleiotropy.



(b) Scenario 2: Multiple mechanisms for the exposure *X*.

### What would happen in each case?

If each diagram is interpretated as a linear structural equations model, we can derive the Wald ratio for each pathway.

| Instruments $Z$              | Pathway $M$           | Effect of $M$ on $X$ | Effect of $M$ on $Y$       | Wald estimand               |  |  |  |  |  |
|------------------------------|-----------------------|----------------------|----------------------------|-----------------------------|--|--|--|--|--|
| Scenario 1                   |                       |                      |                            |                             |  |  |  |  |  |
| $Z_{1,1}, \ldots, Z_{1,p_1}$ | $M_1$                 | $	heta_1$            | $	heta_1eta$               | eta                         |  |  |  |  |  |
| $Z_{2,1}, \ldots, Z_{2,p_2}$ | $M_2$                 | $\theta_2$           | $\theta_2\beta + \alpha_2$ | $\beta + \alpha_2/\theta_2$ |  |  |  |  |  |
| $Z_{3,1},\ldots,Z_{3,p_3}$   | <i>M</i> <sub>3</sub> | $	heta_3$            | $	heta_3eta+lpha_3$        | $eta+lpha_3/	heta_3$        |  |  |  |  |  |
| Scenario 2                   |                       |                      |                            |                             |  |  |  |  |  |
| $Z_{1,1}, \ldots, Z_{1,p_1}$ | $M_1$                 | $	heta_1$            | $	heta_1eta_1$             | $\beta_1$                   |  |  |  |  |  |
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SNPs on the same pathway have the same Wald estimand, while SNPs across different pathways generally have different estimands.

Mechanistic heterogeneity can arise even when all SNPs are valid instruments (Scenario 2).

## Solution 1: Robust likelihood plot



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More detail: Wang, J., Zhao, Q., Bowden, J., Hemani, G., Smith, G. D., Small, D. S., & Zhang, N. R. (2021). Causal inference for heritable phenotypic risk factors using heterogeneous genetic instruments. *PLOS Genetics*. DOI:10.1371/journal.pgen.1009575.

Also contains methods for multiple exposures and overlapping samples.

## Solution 2: Modelling the effect along each path

#### Modified model

► GWAS summary data:

$$\begin{pmatrix} \hat{\gamma}_j \\ \hat{\Gamma}_j \end{pmatrix} \stackrel{ ext{indep.}}{\sim} \mathcal{N}\Big( egin{pmatrix} \gamma_j \\ \beta_j \gamma_j \end{pmatrix}, egin{pmatrix} \sigma_{X_j}^2 & 0 \\ 0 & \sigma_{Y_j}^2 \end{pmatrix} \Big), \quad j=1,\ldots,p,$$

Mixture model for path-specific effects:

$$Z_j \sim ext{Categorical } (\pi_1, \dots, \pi_K),$$
  
 $eta_j | Z_j = k \sim extsf{N}(\mu_k, \sigma_k^2), \quad k = 1, \dots, K.$ 

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More detail: long, D., Zhao, Q., & Chen, Y. (2020). A latent mixture model for heterogeneous causal mechanisms in Mendelian randomization. arXiv:2007.06476.

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- More detail: long, D., Zhao, Q., & Chen, Y. (2020). A latent mixture model for heterogeneous causal mechanisms in Mendelian randomization. arXiv:2007.06476.
- Alternative solution: Bayesian model averaging. See Shapland, C. Y., Zhao, Q., & Bowden, J. (2020). Profile-likelihood Bayesian model averaging for two-sample summary data Mendelian randomization in the presence of horizontal pleiotropy. BioRxiv:2020.02.11.943712.

#### BMI-T2D example: Two-cluster fit



### BMI-T2D example: Posterior intervals



## BMI-T2D example: A possible explanation



### How can we discover the latent pathways? Examine the phenonome!

a1 Graphical representation



#### a2 List representation



cluster list  $\mathbf{L}$ 

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▶ If we let  $\hat{\beta}$  denote the genome-phenome matrix of GWAS coefficients, then  $\hat{\beta}\hat{\beta}^{T}$  should exhibit a low-rank structure.

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#### Additional ideas (Ongoing work)

- ▶ To remove the environmental factors, contrast "signal" loci with "noise" loci.
- ▶ To stablize the results, use "bagging" (bootstrap aggregating).
- ▶ To visualise the results, use lower-dimensional embedding such as the UMAP.

#### Preliminary results: Metabolome GWAS



GWAS data: Kettunen et al. (2016) DOI:10.1038/ncomms11122.

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### Preliminary results: UK BioBank

#### b3 Cluster list

#### Cluster Phenotype Genotype Cluster Phenotype Genotype

| Fat-free                       | Body fat-free mass<br>Standing height<br>Basal metabolic rate | GDF5<br>ZBTB38<br>ID4     | Fat                 | BMI<br>Leg/arm/trunk<br>fat mass           | FTO<br>ADH1B<br>AC09082 |
|--------------------------------|---|---------------------------|---------------------|--|-------------------------|
| Skin/hair<br>color             | Hair color<br>Skin color<br>Ease of skin tanning              | HERC2<br>DEF8<br>TPCN2    | Cardio-<br>vascular | CHD MI<br>High cholesterol<br>Simvastatin  | PXDN<br>SPOCK3          |
| Platelet                       | PLT count<br>Circulatory diseases<br>Heart rate               | JMJD1C<br>CTC-454M9.1     | Diabetes            | Diabetes<br>Metformin<br>Alcohol addiction | TPTE2<br>RP1-116i       |
| Red cell                       | RBC MCH<br>Genetic haemato-<br>logical disorder               | HBS1L<br>ODF3B<br>SLC17A3 | Lymp-<br>hocyte     | Mono count<br>Eos count<br>ANC             | ITGA4<br>CYP8B1<br>GFI1 |
| Venous<br>thrombo-<br>embolism | PE<br>DVT   | ABO<br>SDK1               | Smoking             | Ever smoked<br>Never smoked                | _                       |
| Bone<br>mineral<br>density     | BMD T-score<br>BMD QUI  | SACS<br>FMN2              | Skin<br>neoplasm    | Malignant skin<br>neoplasm                 | PAX5<br>FOXP1           |

## $LMR \times 2$

#### Summary

- Mendelian randomization provides genetic anchors to learn meaningful (and likely causal) representations of life.
- Many challenges remain:
  - 1. Pleiotropy;
  - 2. Non-linear structures and interactions;
  - 3. High-dimensionality;
  - 4. Low signal-to-noise ratio.

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#### Collaborators in the works presented here

Jingshu Wang (Chicago); Dylan Small, Nancy Zhang (UPenn); Gibran Hemani, George Davey Smith (Bristol); Jack Bowden (Exeter); Daniel Iong, Yang Chen (Michigan); Zijun Gao, Trevor Hastie (Stanford).