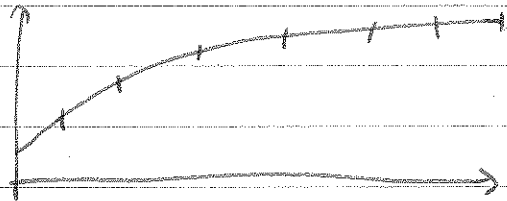


Multifactor IR models

Y_t^i is yield-to-maturity T_i on day t .
 These things are often treated 'as is' from
 a sample MV Gaussian



Warning: Calculate sample cov. matrix V , diagonalize it

$$V \sim \begin{pmatrix} \lambda_1 & & \\ & \ddots & \\ & & \lambda_n \end{pmatrix}$$

Suppose $\lambda_1 > \lambda_2 > \dots > \lambda_n$, $\sum \lambda_i = 1$

We may find $\lambda_1 = 0.95$, $\lambda_2 = 0.02$, $\lambda_3 = 0(10^{-3}) \dots$

So we just need top 2 eigenvalues..?

Taking top 2 components gives 97% of variance.

But the SD of what remains

$$\sqrt{1 - 0.97} = \approx \underline{0.17} ! \text{ not nec. negligible.}$$

We need to consider multifactor models to
 have any hope of capturing yield curve dynamics...

Let us look at a simple MV Gaussian stories.

We propose that

$$r_t = a + b \cdot X_t$$

where X_t is MV O.U process

$$dX_t = \sigma dB_t - A X_t dt$$

A non non-singular and WLOG (by choice
 of a, b) A diagonal (maybe Jordan form)

To solve this, we proceed as in Id:

$$d(\exp(tA)X_t) = e^{tA} dX_t + e^{tA} AX_t dt$$

$$\therefore X_t = e^{-tA} X_0 + \int_0^t e^{(s-t)A} \sigma dB_s$$

and thus

$$\begin{aligned} \int_0^t X_s ds &= (I - e^{-tA})A^{-1} X_0 \\ &\quad + \int_0^t \left(\int_0^s e^{(u-s)A} \sigma dB_u \right) ds \\ &= (I - e^{-tA})A^{-1} X_0 + \int_0^t (I - e^{-(t-u)A})A^{-1} \sigma dB_u \end{aligned}$$

So we have that

$$\int_0^t X_s ds \sim N\left((I - e^{-tA})A^{-1} X_0, V_t \right) \quad \leftarrow \text{known from above}$$

$$\text{and } \int_0^t r_s ds = N\left(at + b \cdot (I - e^{-tA})A^{-1} X_0, b \cdot V_t \cdot b \right)$$

So in this model the ZCB prices are

$$B(0, t) = \exp\left\{ -\left[at + b \cdot (I - e^{-tA})A^{-1} X_0 \right] + \frac{1}{2} b \cdot V_t \cdot b \right\}$$

so the yield to maturity t looks like

$$y(t) = \frac{a + b \cdot (I - e^{-tA})A^{-1} X_0}{t} + \frac{\frac{1}{2} b \cdot V_t \cdot b}{t}$$

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P. 2

How does this look for $N=2$?

$$A = \begin{pmatrix} \lambda & 1 \\ 0 & \lambda \end{pmatrix}$$

We get $\exp tA = e^{\lambda t} \begin{pmatrix} 1 & t \\ 0 & 1 \end{pmatrix}$ so we see a yield curve that looks like

$$y(t) = a + \frac{1 - e^{-\lambda t}}{\lambda t} c_0 + e^{-\lambda t} c_1 + \frac{b \cdot \sqrt{t}}{2t}$$

Nelson - Siegel yield curve "model"

- more flexible, but it is only a fit, not a model.

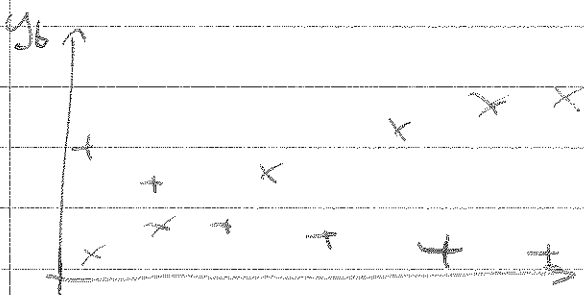
Why do they take A to have Jordan form?

If we had distinct eigenvalues, need to fit

$$y_t = a + \frac{1 - e^{-\lambda_1 t}}{\lambda_1 t} c_1 + \frac{1 - e^{-\lambda_2 t}}{\lambda_2 t} c_2$$

so y_t is a lin. comb. of $e^{-\lambda_i t}$, const.

Now look at data.



x x x full data
+ + + residuals
after taking
out a t

If $\lambda_1 > \lambda_2$, can estimate λ_2 from long-term behaviour, but λ_1 could only be estimated at $t \approx 0 \rightarrow$ not enough data.

Heath-Jarrow-Morton Modelling approach.

The idea here is that we write the bond price in terms of the instantaneous forward rates:

$$B(t, T) = \exp \int_t^T f_{ts} ds$$

and then specify dynamics for process $\{f_{ts} : 0 \leq t \leq s\}$

Note that

$$\lim_{h \rightarrow 0} \frac{1 - B(t, t+h)}{h} = \lim_{h \rightarrow 0} \mathbb{E}_t \left[\frac{1 - e^{-\int_t^{t+h} r_s ds}}{h} \right] = r_t$$

under mild conditions.

So $r_t = f_{tt}$.

There's a myth that HJM and spot rate models are somewhat different, but they are not.

Now let us suppose we want to build a model in a world with just one Brownian motion. Let us suppose that

$$df_{ts} = \sigma_{ts} dB + \alpha_{ts}$$

so we see that

$$f_{ts} = f_{0s} + \int_0^t \sigma_{us} dB_u + \int_0^t \alpha_{us} du$$

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and so
$$\int_t^T f_{ts} = \int_t^T \theta_{fs} + \int_0^t \left(\int_t^T \sigma_{vs} ds \right) dB_v$$

$$+ \int_0^t \left(\int_t^T \alpha_{vs} ds \right) dv$$

Now notice that for T fixed

$$M(t, T) = E_t \left[\exp \left(- \int_0^T r_s ds \right) \right] = \exp \left(- \int_0^t r_s ds \right) B(t, T)$$

$$= \exp \left(- \int_0^t \sum_{ST} dB_s - \frac{1}{2} \int_0^t \sum_{ST}^2 ds \right)$$

$$\underbrace{\exp \left(- \int_0^T f_{0s} ds \right)}_{B(0, T)}$$

