

# When to efficiently rebalance a portfolio

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## Continuous-time finance

Denote by  $S = (S^1, \dots, S^d)$  the price process vector of  $d$  risky assets. Under a risk-free rate  $r$ , the wealth process  $V$  for an investment strategy  $H = (H^1, \dots, H^d)$  and a consumption plan  $c$  is described as

$$dV_t = \sum_{i=0}^d H_t^i dS_t^i - c_t dt,$$

where

$$\frac{dS_t^0}{S_t^0} = r_t dt, \quad H^0 = \frac{1}{S^0} \left( V - \sum_{i=1}^d H^i S^i \right).$$

Under an admissibility condition  $V > 0$ , this can be rewritten as

$$\frac{dV_t}{V_t} = \sum_{i=0}^d \pi_t^i \frac{dS_t^i}{S_t^i} - \frac{c_t}{V_t} dt, \quad \sum_{i=0}^d \pi^i = 1$$

in terms of  $\pi^i = H^i S^i / V$

# A constant weight asset allocation

Here,  $\pi^i = H^i S^i / V$  represents the fraction of the wealth invested in the  $i$ th asset.

In this study we will assume  $\pi = (\pi^1, \dots, \pi^d)$  is **continuous** and of **finite variation**. This includes the case it is a constant, that is, a **constant weight asset allocation**.

- Kelly criterion
- CRRA utility maximization (Merton strategy)
- Epstein-Zin utility maximization
- Relative performance criteria
- $1/N$  strategy

## A concrete example

Assume

$$\frac{dS_t^i}{S_t^i} = \mu_t^i dt + \sigma_t^i dW_t^i, \quad \sigma_t^i \sigma_t^j d\langle W^i, W^j \rangle_t = \Sigma_t^{ij} dt$$

for adapted processes  $\mu^i$  and  $\Sigma^{ij}$  with  $\Sigma = [\Sigma^{ij}]$  regular. Under

$$\frac{dV_t}{V_t} = \sum_{i=0}^d \pi_t^i \frac{dS_t^i}{S_t^i},$$

Itô's formula gives

$$\mathbb{E} \left[ \log \frac{V_T}{V_0} \right] = \int_0^T \mathbb{E} \left[ -\frac{1}{2} (\pi_t - \theta_t)^\top \Sigma_t (\pi_t - \theta_t) + \frac{1}{2} \theta_t^\top \Sigma_t \theta_t \right] dt,$$

where

$$\theta := \Sigma^{-1}(\mu - r\mathbf{1}), \quad \mu = (\mu^1, \dots, \mu^d)^\top, \quad \mathbf{1} = (1, \dots, 1)^\top.$$

The growth optimal portfolio is therefore  $\pi = \theta$ .

## Objective

Notice  $H^i = V\pi^i/S^i$  is not of finite variation, needs to be approximated by a simple predictable process  $H^{n,i}$ . Denote by

$$Z_t^n = V_t - V_t^n = \sum_{i=0}^n \int_0^t (H_u^i - H_u^{n,i}) dS_u^i$$

the tracking error process,. We want to minimize

$$E[\langle Z^n \rangle_T] = \sum_{i,j=1}^d E \left[ \int_0^T (H_u^i - H_u^{n,i})(H_u^j - H_u^{n,j}) d\langle S^i, S^j \rangle_u \right]$$

under a cost constraint

$$E[C_T^n] \leq c_n,$$

where  $C^n$  is the counting process of the jumps of the  $d$  dimensional simple predictable process  $(H^{n,1}, \dots, H^{n,d})$ .

## Note on the objective

On the tracking error,

- $E[\langle Z^n \rangle_T] = E[|Z_T^n|^2]$  under any martingale measure.
- $E[\langle Z^n \rangle_T]$  is time consistent, while  $E[|Z_T^n|^2]$  is not in general.
- it is the cumulative local mean squared error (c.f. local mean variance criteria):

$$\sum_{i=1}^N E[|Z_{t_i}^n - Z_{t_{i-1}}^n|^2] \approx E[\langle Z^n \rangle_T],$$

where  $0 = t_0 < \dots < t_N = T$ .

The cost is

- the expected number of transactions.
- the expected cumulative fixed transaction costs.

Variational inequality approach under the Black-Scholes model:

- Martini and Patry (1999) for mean squared error

## Fixed transaction costs under the Black-Scholes model

- Schroder (1995), Whalley-Wilmott (1997), Korn (1998), Liu(2004), Lo et al. (2004) : for CARA utility
- Morton and Pliska (1995), Atkinson and Wilmott (1995) : growth optimal portfolio under quasi-fixed transaction costs
- Altrarovici, Muhle-Karbe and Soner (2015): asymptotics for CRRA; No-trade region is defined as

$$\left\{ \pi; (\pi - \pi^*)^\top M(\pi - \pi^*) \leq h \right\},$$

where  $\pi^*$  is the frictionless Merton proportion and  $M$  is the solution of “a matrix valued algebraic Riccati equation”. Asymptotically optimal is “the strategy that consumes at the frictionless Merton rate, does not trade while the current position lies in the above no-trade region, and jumps to the frictionless Merton proportion once its boundaries are breached.”

# High-frequency asymptotic approach: one-dimensional case

Fukasawa (2011, 2014): For general one-dimensional continuous semimartingales  $H$  and  $S$  (corresponding to the case  $d = 1$ ),

$$\lim_{n \rightarrow \infty} E[C_T^n] E[\langle Z^n \rangle_T] \geq \frac{1}{6} E\left[\int_0^T K_t d\langle H \rangle_t\right]^2$$

for “any” sequence of simple predictable processes with  $E[C_T^n] \rightarrow \infty$ , under the condition

$$d\langle S \rangle = K_t^2 d\langle H \rangle_t$$

with continuous  $K$  (without assuming  $\pi$  is of finite variation). Equivalently,

$$E[\langle Z^n \rangle_T] \geq \frac{1}{6n} E\left[\int_0^T K_t d\langle H \rangle_t\right]^2 + o\left(\frac{1}{n}\right), \quad n \rightarrow \infty$$

under  $E[C_T^n] \leq n$ . The lower bound is attained by discretizations of  $H$  with

$$\tau_{j+1}^n = \inf \left\{ t > \tau_j^n; K_{\tau_j^n} |H_t - H_{\tau_j^n}|^2 = h_n \right\}, \quad h_n \rightarrow 0.$$

Related works: Fukasawa (2011), Cai et al. (2016, 2018).



## Multi-dimensional case

Gobet and Landon (2014), Gobet and Stazhynski (2018): extension to  $d$  dimensional cases when the strategy is of the form  $H_t^i = f^i(t, S_t, R_t)$  with  $R$  of finite variation.

- in the context of hedging
- approximations by discretizations

Our contribution:

- investment strategies
- optimality among simple predictable strategies

### Our structure condition

The strategy  $H = (H^1, \dots, H^d)$  has a continuous drift and there exist regular matrix valued continuous adapted processes  $J$  and  $K$  such that

$$d\langle H, H \rangle_t = J_t dt, \quad d\langle S, S \rangle_t = K_t^\top J_t K_t dt.$$

# On the structure condition

## Theorem A

Let  $\pi^0$  and  $\pi = (\pi^1, \dots, \pi^d)$  be continuous adapted processes of finite variation with  $\sum_{i=0}^d \pi^i = 1$  and assume

$$\frac{dV_t}{V_t} = \sum_{i=0}^d \pi_t^i \frac{dS_t^i}{S_t^i} - c_t dt,$$

$$d\langle \log S^i, \log S^j \rangle_t = \Sigma_t^{ij} dt, \quad d \log S_t^0 = r_t dt$$

for continuous adapted processes  $c$ ,  $r$ ,  $\mu^i$  (the drift of  $S^i$ ) and  $\Sigma^{ij}$ . If  $\Sigma_t = [\Sigma_t^{ij}]$  is regular and  $\pi_t^i \neq 0$  for all  $t$  and  $i \geq 0$ , then the structure condition is met with

$$J = G^\top G, \quad G = \Sigma^{1/2} (H^\top \pi - \text{diag}(H)), \quad K = J^{-1/2} \Sigma^{1/2} \text{diag}(S)$$

for  $H = (H^1, \dots, H^d)$ ,  $H^i = V \pi^i / S^i$ .

## Admissible sequence

For a given strategy  $H = (H^1, \dots, H^d)$  with the structure condition, we say a sequence of  $d$  dimensional simple predictable processes

$$(H^{n,1}, \dots, H^{n,d}), \quad H_t^{n,i} = \sum_{j=0}^{\infty} H_j^{n,i} 1_{(\tau_j^n, \tau_{j+1}^n]}(t),$$

where  $\{\tau_j^n\}$  is a sequence of  $\{\mathcal{F}_t\}$ -stopping times and  $H_j^{n,i}$  is  $\mathcal{F}_{\tau_j^n}$ -measurable, is **admissible** for  $H$  if

- 1  $\sup_{t \in [0, T]} |H_t^i - H_t^{n,i}| \rightarrow 0$  in probability for each  $i = 1, \dots, d$ , and
- 2  $E[C_T^n \langle Z^n \rangle_T]$  is uniformly integrable, where  $C_T^n = \max\{j; \tau_j^n \leq T\}$  and

$$\langle Z^n \rangle_T = \sum_{i,j=1}^d \int_0^T (H_u^i - H_u^{n,i})(H_u^j - H_u^{n,j}) d\langle S^i, S^j \rangle_u.$$

We do **NOT** assume a priori that  $H_j^{n,i} = H_{\tau_j^n}^i$ .

# Main result

## Theorem B

Under the structure condition, for any admissible sequence,

$$\lim_{n \rightarrow \infty} \mathbb{E}[C_T^n] \mathbb{E}[\langle Z^n \rangle_T] \geq \mathbb{E} \left[ \int_0^T \text{Tr}(U_t J_t) dt \right]^2,$$

where  $U_t = u(J_t, K_t)$  with a map  $u$  defined later. Further, if  $H_j^{n,i} = H_{\tau_j^n}^i$  with

$$\tau_{j+1}^n = \inf \left\{ t > \tau_j^n; \left( H_t - H_{\tau_j^n} \right)^\top U_{\tau_j^n} \left( H_t - H_{\tau_j^n} \right) \geq h_n \right\},$$

then there exists a sequence of stopping times  $\sigma^m \uparrow T$  such that

$$\lim_{n \rightarrow \infty} \mathbb{E}[C_{\sigma^m}^n] \mathbb{E}[\langle Z^n \rangle_{\sigma^m}] = \mathbb{E} \left[ \int_0^{\sigma^m} \text{Tr}(U_t J_t) dt \right]^2.$$

# A matrix-valued algebraic Riccati equation

- Atkinson and Wilmott (1995)
- Gobet and Landon (2014)
- Altarovici, Muhle-Karbe and Soner (2015)

An extension of Lemma from Gobet and Landon (2014):

## Lemma

Denote by  $\mathcal{M}_d$  and  $\mathcal{S}_d$  respectively the sets of  $d \times d$  matrices and positive definite matrices. For any  $J \in \mathcal{S}_d$  and  $K \in \mathcal{M}_d$ , there exists a unique nonnegative definite matrix  $U = u(J, K)$  such that

$$2\text{Tr}(JU)U + 4UJU = K^\top JK$$

and the map  $u$  is continuous on  $\mathcal{M}_d \times \mathcal{S}_d$ . Further, if  $K^\top JK \in \mathcal{S}_d$ , then  $U = u(J, K) \in \mathcal{S}_d$ .

## Towards the proof: one dimensional case revisited

When  $d = 1$ ,

$$\langle Z^n \rangle_T = \sum_{j \geq 0} \int_{\tau_j^n \wedge T}^{\tau_{j+1}^n \wedge T} (H_t - H_j^n)^2 K_t^2 d\langle H \rangle_t$$

and by Itô's formula,

$$\begin{aligned} \int_{\tau_j^n}^{\tau_{j+1}^n} (H_t - H_j^n)^2 K_{\tau_j^n}^2 d\langle H \rangle_t &= \frac{1}{6} K_{\tau_j^n}^2 \left( (H_{\tau_{j+1}^n} - H_j^n)^4 - (H_{\tau_j^n} - H_j^n)^4 \right) \\ &\quad - \frac{2}{3} K_{\tau_j^n}^2 \int_{\tau_j^n}^{\tau_{j+1}^n} (H_t - H_j^n)^3 dH_t. \end{aligned}$$

We have

$$\mathbb{E}[\langle Z^n \rangle_T] \approx \frac{1}{6} \mathbb{E} \left[ \sum K_{\tau_j^n}^2 \mathbb{E}[\left( (\Delta_j + \delta_j)^4 - \delta_j^4 \right) | \mathcal{F}_{\tau_j^n}] \right]$$

for  $\Delta_j = H_{\tau_{j+1}^n} - H_{\tau_j^n}$  and  $\delta_j = H_{\tau_j^n} - H_j^n$ .

## Pearson's inequality

Let  $\Delta$  be a centered random variable. Then,

$$E[\Delta^4]E[\Delta^2] \geq E[\Delta^3]^2 + E[\Delta^2]^3.$$

One way to prove this is to notice that for any  $(a, b)$ ,

$$\begin{aligned} 0 &\leq E[(a(\Delta^2 - E[\Delta^2]) + b\Delta)^2] \\ &= (a, b) \begin{pmatrix} E[\Delta^4] - E[\Delta^2]^2 & E[\Delta^3] \\ E[\Delta^3] & E[\Delta^2] \end{pmatrix} \begin{pmatrix} a \\ b \end{pmatrix}. \end{aligned}$$

To apply this to our problem (assuming  $E[\Delta_j | \mathcal{F}_{\tau_j^n}] = 0$ ), note

$$\begin{aligned} E[((\Delta_j + \delta_j)^4 - \delta_j^4) | \mathcal{F}_{\tau_j^n}] &= E[\Delta_j^4 | \mathcal{F}_{\tau_j^n}] + 4\delta_j E[\Delta_j^3 | \mathcal{F}_{\tau_j^n}] + 6\delta_j^2 E[\Delta_j^2 | \mathcal{F}_{\tau_j^n}] \\ &= 6E[\Delta_j^2 | \mathcal{F}_{\tau_j^n}] \left( \delta_j + \frac{1}{3} \frac{E[\Delta_j^3 | \mathcal{F}_{\tau_j^n}]}{E[\Delta_j^2 | \mathcal{F}_{\tau_j^n}]} \right)^2 + E[\Delta_j^4 | \mathcal{F}_{\tau_j^n}] - \frac{2}{3} \frac{E[\Delta_j^3 | \mathcal{F}_{\tau_j^n}]^2}{E[\Delta_j^2 | \mathcal{F}_{\tau_j^n}]}. \end{aligned}$$

## Multidimensional case

For  $U = u(J, K)$ , the solution of the matrix-valued algebraic Riccati equation

$$2\text{Tr}(JU)U + 4UJU = K^\top JK,$$

by Itô's formula applied to the  $C^2$  function  $f(x) = (x^\top Ux)^2$ , we have

$$\begin{aligned} E[\langle Z^n \rangle_T] &= \sum_{i,j=1}^d E\left[\int_0^T (H_t^i - H_t^{n,i})(H_t^j - H_t^{n,j})K_t^i K_t^j d\langle H^i, H^j \rangle_t\right] \\ &\approx E\left[\sum \left((\Delta_j + \delta_j)^\top (\Delta_j + \delta_j)\right)^2 - \left(\delta_j^\top \delta_j\right)^2\right] \end{aligned}$$

for

$$\Delta_j = U_{\tau_j^n}^{1/2} (H_{\tau_{j+1}^n}^1 - H_{\tau_j^n}^1, \dots, H_{\tau_{j+1}^n}^d - H_{\tau_j^n}^d)^\top,$$

$$\delta_j = U_{\tau_j^n}^{1/2} (H_{\tau_j^n}^1 - H_j^{n,1}, \dots, H_{\tau_j^n}^d - H_j^{n,d})^\top.$$



Let  $\Delta$  be a  $d$ -dimensional  $L^4$  random variable with  $E[\Delta] = 0$  and  $\delta \in \mathbb{R}^d$ .

$$\begin{aligned} & E[((\Delta + \delta)^T (\Delta + \delta))^2] - (\delta^T \delta)^2 \\ &= E[(\Delta^T \Delta + 2\delta^T \Delta + \delta^T \delta)^2] - (\delta^T \delta)^2 \\ &= E[(\Delta^T \Delta)^2] + 4\delta^T E[\Delta \Delta^T] \delta + 4E[\delta^T \Delta (\Delta^T \Delta)] + 2\delta^T \delta E[\Delta^T \Delta]. \end{aligned}$$

Taking the gradient with respect to  $\delta$ ,

$$2(4E[\Delta \Delta^T]) + 2E[\Delta^T \Delta])\delta + 4E[\Delta (\Delta^T \Delta)]$$

and so, the minimum is attained at

$$\delta = -(2E[\Delta \Delta^T] + E[\Delta^T \Delta]I)^{-1}E[\Delta (\Delta^T \Delta)].$$

Substitute this to get

$$\begin{aligned} & E[((\Delta + \delta)^T (\Delta + \delta))^2] - (\delta^T \delta)^2 \\ & \geq E[(\Delta^T \Delta)^2] - E[\Delta (\Delta^T \Delta)] \left( E[\Delta \Delta^T] + \frac{1}{2}E[\Delta^T \Delta]I \right)^{-1} E[\Delta (\Delta^T \Delta)]. \end{aligned}$$

# A multi-dimensional extension of Pearson's inequality

## Theorem C

Let  $\Delta$  be a  $d$ -dimensional  $L^4$  random variable with  $E[\Delta] = 0$  and  $\delta \in \mathbb{R}^d$ . Then, for any positive definite matrix  $A$ ,

$$E[(\Delta^T \Delta)^2] - E[\Delta(\Delta^T \Delta)] \left( E[\Delta \Delta^T] + A \right)^{-1} E[\Delta(\Delta^T \Delta)] \geq E[\Delta^T \Delta]^2.$$

Proof: for any  $\alpha \in \mathbb{R}$  and  $\beta \in \mathbb{R}^d$ ,

$$E[(\alpha(\Delta^T \Delta - E[\Delta^T \Delta]) + \beta^T \Delta)^2] \geq 0$$

The left hand side is a quadratic form with respect to the symmetric matrix

$$\begin{pmatrix} E[(\Delta^T \Delta - E[\Delta^T \Delta])^2] & E[\Delta^T (\Delta^T \Delta)] \\ E[\Delta(\Delta^T \Delta)] & E[\Delta \Delta^T] \end{pmatrix}$$

and the above nonnegativity implies that the matrix is nonnegative definite.

(continued) Therefore the matrix

$$\begin{pmatrix} E[(\Delta^T \Delta - E[\Delta^T \Delta])^2] & E[\Delta^T (\Delta^T \Delta)] \\ E[\Delta (\Delta^T \Delta)] & E[\Delta \Delta^T] + A \end{pmatrix}$$

is a positive definite and so, has a positive determinant. By the determinant formula for block matrices, the determinant is computed as

$$\begin{aligned} & \left| E[\Delta \Delta^T] + A \right| \\ & \times \left( E[(\Delta^T \Delta - E[\Delta^T \Delta])^2] - E[\Delta (\Delta^T \Delta)] \left( E[\Delta \Delta^T] + A \right)^{-1} E[\Delta (\Delta^T \Delta)] \right). \end{aligned}$$

# Summary

- We can asymptotically efficiently discretize a rebalancing strategy.
- The rebalancing times should be hitting times of an ellipsoid.
- The ellipsoid is stochastic, the solution of a quadratic equation with the covariance matrices.
- Discretization is an efficient approximation, for which the proof requires a nontrivial multidimensional extension of Pearson's inequality.