## Paper 1, Section I

7H Statistics

Consider an estimator  $\hat{\theta}$  of an unknown parameter  $\theta$ , and assume that  $\mathbb{E}_{\theta}(\hat{\theta}^2) < \infty$  for all  $\theta$ . Define the *bias* and *mean squared error* of  $\hat{\theta}$ .

Show that the mean squared error of  $\hat{\theta}$  is the sum of its variance and the square of its bias.

Suppose that  $X_1, \ldots, X_n$  are independent identically distributed random variables with mean  $\theta$  and variance  $\theta^2$ , and consider estimators of  $\theta$  of the form  $k\bar{X}$  where  $\bar{X} = \frac{1}{n} \sum_{i=1}^{n} X_i$ .

- (i) Find the value of k that gives an unbiased estimator, and show that the mean squared error of this unbiased estimator is  $\theta^2/n$ .
- (ii) Find the range of values of k for which the mean squared error of  $k\bar{X}$  is smaller than  $\theta^2/n$ .

### Paper 2, Section I

#### 8H Statistics

There are 100 patients taking part in a trial of a new surgical procedure for a particular medical condition. Of these, 50 patients are randomly selected to receive the new procedure and the remaining 50 receive the old procedure. Six months later, a doctor assesses whether or not each patient has fully recovered. The results are shown below:

	Fully	Not fully
	recovered	recovered
Old procedure	25	25
New procedure	31	19

The doctor is interested in whether there is a difference in full recovery rates for patients receiving the two procedures. Carry out an appropriate 5% significance level test, stating your hypotheses carefully. [You do not need to derive the test.] What conclusion should be reported to the doctor?

[Hint: Let  $\chi_k^2(\alpha)$  denote the upper 100 $\alpha$  percentage point of a  $\chi_k^2$  distribution. Then

$$\chi_1^2(0.05) = 3.84, \ \chi_2^2(0.05) = 5.99, \ \chi_3^2(0.05) = 7.82, \ \chi_4^2(0.05) = 9.49.$$

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Paper 4, Section II

**19H Statistics** 

Consider a linear model

$$\mathbf{Y} = X\boldsymbol{\beta} + \boldsymbol{\varepsilon},\tag{\dagger}$$

where X is a known  $n \times p$  matrix,  $\beta$  is a  $p \times 1$  (p < n) vector of unknown parameters and  $\varepsilon$  is an  $n \times 1$  vector of independent  $N(0, \sigma^2)$  random variables with  $\sigma^2$  unknown. Assume that X has full rank p. Find the least squares estimator  $\hat{\beta}$  of  $\beta$  and derive its distribution. Define the residual sum of squares RSS and write down an unbiased estimator  $\hat{\sigma}^2$  of  $\sigma^2$ .

Suppose that  $V_i = a + bu_i + \delta_i$  and  $Z_i = c + dw_i + \eta_i$ , for  $i = 1, \ldots, m$ , where  $u_i$  and  $w_i$  are known with  $\sum_{i=1}^m u_i = \sum_{i=1}^m w_i = 0$ , and  $\delta_1, \ldots, \delta_m, \eta_1, \ldots, \eta_m$  are independent  $N(0, \sigma^2)$  random variables. Assume that at least two of the  $u_i$  are distinct and at least two of the  $w_i$  are distinct. Show that  $\mathbf{Y} = (V_1, \ldots, V_m, Z_1, \ldots, Z_m)^T$  (where T denotes transpose) may be written as in  $(\dagger)$  and identify X and  $\boldsymbol{\beta}$ . Find  $\hat{\boldsymbol{\beta}}$  in terms of the  $V_i, Z_i, u_i$  and  $w_i$ . Find the distribution of  $\hat{b} - \hat{d}$  and derive a 95% confidence interval for b - d.

[Hint: You may assume that  $\frac{RSS}{\sigma^2}$  has a  $\chi^2_{n-p}$  distribution, and that  $\hat{\beta}$  and the residual sum of squares are independent. Properties of  $\chi^2$  distributions may be used without proof.]

## Paper 1, Section II

### **19H** Statistics

Suppose that  $X_1$ ,  $X_2$ , and  $X_3$  are independent identically distributed Poisson random variables with expectation  $\theta$ , so that

$$\mathbb{P}(X_i = x) = \frac{e^{-\theta}\theta^x}{x!} \quad x = 0, 1, \dots,$$

and consider testing  $H_0: \theta = 1$  against  $H_1: \theta = \theta_1$ , where  $\theta_1$  is a known value greater than 1. Show that the test with critical region  $\{(x_1, x_2, x_3): \sum_{i=1}^3 x_i > 5\}$  is a likelihood ratio test of  $H_0$  against  $H_1$ . What is the size of this test? Write down an expression for its power.

A scientist counts the number of bird territories in n randomly selected sections of a large park. Let  $Y_i$  be the number of bird territories in the *i*th section, and suppose that  $Y_1, \ldots, Y_n$  are independent Poisson random variables with expectations  $\theta_1, \ldots, \theta_n$  respectively. Let  $a_i$  be the area of the *i*th section. Suppose that n = 2m,  $a_1 = \cdots = a_m = a(> 0)$  and  $a_{m+1} = \cdots = a_{2m} = 2a$ . Derive the generalised likelihood ratio  $\Lambda$  for testing

$$H_0: \theta_i = \lambda a_i \text{ against } H_1: \theta_i = \begin{cases} \lambda_1 & i = 1, \dots, m \\ \lambda_2 & i = m+1, \dots, 2m. \end{cases}$$

What should the scientist conclude about the number of bird territories if  $2\log_e(\Lambda)$  is 15.67?

[*Hint:* Let  $F_{\theta}(x)$  be  $\mathbb{P}(W \leq x)$  where W has a Poisson distribution with expectation  $\theta$ . Then

 $F_1(3) = 0.998$ ,  $F_3(5) = 0.916$ ,  $F_3(6) = 0.966$ ,  $F_5(3) = 0.433$ .

## Paper 3, Section II

20H Statistics

Suppose that  $X_1, \ldots, X_n$  are independent identically distributed random variables with

$$\mathbb{P}(X_i = x) = \binom{k}{x} \theta^x (1-\theta)^{k-x}, \quad x = 0, \dots, k,$$

where k is known and  $\theta$  (0 <  $\theta$  < 1) is an unknown parameter. Find the maximum likelihood estimator  $\hat{\theta}$  of  $\theta$ .

Statistician 1 has prior density for  $\theta$  given by  $\pi_1(\theta) = \alpha \theta^{\alpha-1}$ ,  $0 < \theta < 1$ , where  $\alpha > 1$ . Find the posterior distribution for  $\theta$  after observing data  $X_1 = x_1, \ldots, X_n = x_n$ . Write down the posterior mean  $\hat{\theta}_1^{(B)}$ , and show that

$$\hat{\theta}_1^{(B)} = c\,\hat{\theta} + (1-c)\tilde{\theta}_1,$$

where  $\tilde{\theta}_1$  depends only on the prior distribution and c is a constant in (0, 1) that is to be specified.

Statistician 2 has prior density for  $\theta$  given by  $\pi_2(\theta) = \alpha(1-\theta)^{\alpha-1}$ ,  $0 < \theta < 1$ . Briefly describe the prior beliefs that the two statisticians hold about  $\theta$ . Find the posterior mean  $\hat{\theta}_2^{(B)}$  and show that  $\hat{\theta}_2^{(B)} < \hat{\theta}_1^{(B)}$ .

Suppose that  $\alpha$  increases (but n, k and the  $x_i$  remain unchanged). How do the prior beliefs of the two statisticians change? How does c vary? Explain briefly what happens to  $\hat{\theta}_1^{(B)}$  and  $\hat{\theta}_2^{(B)}$ .

[*Hint:* The  $Beta(\alpha, \beta)$  ( $\alpha > 0, \beta > 0$ ) distribution has density

$$f(x) = \frac{\Gamma(\alpha + \beta)}{\Gamma(\alpha)\Gamma(\beta)} x^{\alpha - 1} (1 - x)^{\beta - 1}, \quad 0 < x < 1,$$

with expectation  $\frac{\alpha}{\alpha+\beta}$  and variance  $\frac{\alpha\beta}{(\alpha+\beta+1)(\alpha+\beta)^2}$ . Here,  $\Gamma(\alpha) = \int_0^\infty x^{\alpha-1} e^{-x} dx$ ,  $\alpha > 0$ , is the Gamma function.]